

# Forecasting of stationary and ARIMA processes \*

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# 1. Wiener-Kolmogorov formula

According to Wold's decomposition theorem, a second-order stationary process (with mean zero) can be written in the following form:

$$X_t = Y_t + D_t$$

where

$$Y_t = \sum_{j=0}^{\infty} d_j u_{t-j}, \quad d_0 = 1, \quad \{u_t\} \sim BB(0, \sigma^2),$$

$D_t$  is deterministic,

$$u_t = X_t - P_L(X_t | X_{t-1}, X_{t-2}, \dots).$$

Consequently,

$$P_L(u_t | X_{t-1}, X_{t-2}, \dots) = 0$$

or, more generally,

$$P_L(u_t | X_{t-\ell}, X_{t-\ell-1}, \dots) = 0, \quad \forall \ell \geq 1.$$

If  $X_t$  is a strictly indeterministic process,

$$X_t = \sum_{j=0}^{\infty} d_j u_{t-j},$$

we have:

$$\begin{aligned} P_L(X_t | X_{t-1}, X_{t-2}, \dots) &= P_L(u_t | X_{t-1}, X_{t-2}, \dots) \\ &+ P_L \left[ \sum_{j=1}^{\infty} d_j u_{t-j} | X_{t-1}, X_{t-2}, \dots \right] \\ &= \sum_{j=1}^{\infty} d_j u_{t-j}. \end{aligned}$$

If we furthermore suppose that the  $\{u_t\}$  are independent,

$$\begin{aligned} E(u_t | X_{t-1}, X_{t-2}, \dots) &= E(u_t | u_{t-1}, u_{t-2}, \dots) = 0, \\ E(X_t | X_{t-1}, X_{t-2}, \dots) &= \sum_{j=1}^{\infty} d_j u_{t-j} \end{aligned}$$

$$= P_L(X_t | X_{t-1}, X_{t-2}, \dots) .$$

We have the best prediction in the mean-square-error (MSE) sense.

Let  $\{X_t\}$  a weakly stationary indeterministic process:

$$X_t = \sum_{j=0}^{\infty} d_j u_{t-j} = d(B) u_t, \quad (1.1)$$

where

$$d_0 = 1, \quad d(B) = \sum_{j=0}^{\infty} d_j B^j . \quad (1.2)$$

Let us denote:

$$P_{t-j} X_t = P(X_t | X_{t-j}, X_{t-j-1}, \dots) .$$

Then,

$$\begin{aligned} P_{t-j} X_t &= X_t \text{ for } j \leq 0, \\ u_t &= X_t - P(X_t | X_{t-1}, X_{t-2}, \dots) \\ &= X_t - P_{t-1} X_t \\ P_{t-1} X_t &= \sum_{j=0}^{\infty} d_j P_{t-1} u_{t-j} = \sum_{j=1}^{\infty} d_j u_{t-j} \\ &= \left( \frac{d(B)}{B} \right)_+ u_{t-1} \end{aligned}$$

where we define

$$\begin{aligned} \left( \sum_{j=-\infty}^{+\infty} h_j B^j \right)_+ &= \sum_{j=0}^{\infty} h_j B^j, \\ \left[ \frac{d(B)}{B} \right]_+ &= (d_0 B^{-1} + d_1 B^0 + d_2 B^1 + \dots) \\ &= \left( \sum_{j=0}^{\infty} d_j B^{j-1} \right)_+ \\ &= \sum_{j=1}^{\infty} d_j B^{j-1}. \end{aligned} \quad (1.3)$$

Similarly, we get at lag  $\ell$ ,

$$\begin{aligned} P_{t-\ell}X_t &= \sum_{j=0}^{\infty} d_j P_{t-\ell} u_{t-j} = \sum_{j=\ell}^{\infty} d_j u_{t-j} \\ &= \left( \frac{d(B)}{B^\ell} \right)_+ u_{t-\ell} \end{aligned}$$

or, equivalently,

$$P_t X_{t+\ell} = \left( \frac{d(B)}{B^\ell} \right)_+ u_t = \sum_{j=\ell}^{\infty} d_j u_{t+\ell-j}. \quad (1.4)$$

The formula (1.4) is called the Wiener-Kolmogorov formula for linear prediction  $\ell$  periods ahead. We see that  $P_t X_{t+\ell}$  can be computed by dropping the most recent  $\ell$  terms from the Wold decomposition:

$$\begin{aligned} X_t &= \sum_{j=0}^{\infty} d_j u_{t-j} \\ &= \sum_{j=0}^{\ell-1} d_j u_{t+\ell-j} + \sum_{j=\ell}^{\infty} d_j u_{t+\ell-j} \\ &= e_t(\ell) + P_t X_{t+\ell}. \end{aligned} \quad (1.5)$$

If  $X_t$  is invertible, we can write

$$u_t = \frac{1}{d(B)} X_t$$

hence

$$\begin{aligned} P_{t-\ell} X_t &= \left( \frac{d(B)}{B^\ell} \right)_+ \frac{1}{d(B)} X_{t-\ell}, \\ P_t X_{t+\ell} &= \left( \frac{d(B)}{B^\ell} \right)_+ \frac{1}{d(B)} X_t, \quad \ell \geq 1. \end{aligned}$$

**1.1 Example** For an AR(1) process,

$$X_t = \varphi_1 X_{t-1} + u_t, \quad |\varphi_1| < 1 \quad (1.6)$$

we have the representation

$$X_t = \frac{1}{1 - \varphi_1 B} u_t, \quad (1.7)$$

hence

$$\begin{aligned} P_{t-1}X_t &= \varphi_1 P_{t-1}X_{t-1} + P_{t-1}u_t \\ &= \varphi_1 X_{t-1}, \end{aligned} \tag{1.8}$$

$$\begin{aligned} P_{t-\ell}X_t &= \varphi_1 P_{t-\ell}X_{t-1} \\ &= \varphi_1 P_{t-\ell}(\varphi_1 X_{t-2} + u_{t-1}) \\ &= \varphi_1^2 P_{t-\ell}X_{t-2} \\ &= \varphi_1^\ell X_{t-\ell}. \end{aligned}$$

If we use the Wiener-Kolmogorov formula, we get:

$$\begin{aligned} P_{t-\ell}X_t &= \left[ B^{-\ell} \frac{1}{1 - \varphi_1 B} \right]_+ u_{t-\ell} \\ &= [B^{-\ell} (1 + \varphi_1 B + \varphi_1^2 B^2 + \dots)]_+ (1 - \varphi_1 B) X_{t-\ell} \\ &= \varphi_1^\ell (1 + \varphi_1 B + \varphi_1^2 B^2 + \dots) (1 - \varphi_1 B) X_{t-\ell} \\ &= \frac{\varphi_1^\ell}{1 - \varphi_1 B} (1 - \varphi_1 B) X_{t-\ell} \\ &= \varphi_1^\ell X_{t-\ell}. \end{aligned}$$

**1.2 Example** For an MA(1) process,

$$X_t = (1 - \theta_1 B) u_t, \quad |\theta_1| < 1, \tag{1.9}$$

we have

$$u_t = \frac{1}{1 - \theta_1 B} X_t, \tag{1.10}$$

hence the following forecasts: for  $\ell = 1$ ,

$$\begin{aligned} P_{t-1}X_t &= [B^{-1} (1 - \theta_1 B)]_+ \frac{1}{1 - \theta_1 B} X_{t-1} \\ &= \frac{-\theta_1}{1 - \theta_1 B} X_{t-1} \\ &= -\theta_1 \sum_{i=0}^{\infty} \theta_1^i X_{t-1-i} \\ &= -\sum_{i=1}^{\infty} \theta_1^i X_{t-i}; \end{aligned} \tag{1.11}$$

pour  $\ell \geq 2$

$$P_{t-\ell}X_t = [B^{-\ell}(1 - \theta_1 B)]_+ \frac{1}{1 - \theta_1 B} X_{t-\ell} = 0.$$

**1.3 Example** For an ARMA(1, 1) process of the form

$$\begin{aligned} (1 - \varphi_1 B) X_t &= (1 - \theta_1 B) u_t, \\ |\varphi_1| < 1, \quad |\theta_1| < 1, \end{aligned} \tag{1.12}$$

the Wold representation can be written:

$$\begin{aligned} X_t &= \frac{(1 - \theta_1 B)}{(1 - \varphi_1 B)} u_t \\ &= d(B) u_t. \end{aligned} \tag{1.13}$$

From the latter, we then get the following forecasts: at the horizon 1,

$$P_{t-1}X_t = \left[ \frac{d(B)}{B} \right]_+ u_{t-1} = \left[ \frac{d(B)}{B} \right]_+ \frac{1}{d(B)} X_{t-1} \tag{1.14}$$

where

$$\begin{aligned} \left[ \frac{d(B)}{B} \right]_+ &= \left[ B^{-1} \frac{(1 - \theta_1 B)}{(1 - \varphi_1 B)} \right]_+ \\ &= \left[ \frac{B^{-1}}{1 - \varphi_1 B} - \frac{\theta_1}{1 - \varphi_1 B} \right]_+ \\ &= \left[ B^{-1} [1 + \varphi_1 B (1 + \varphi_1 B + \varphi_1^2 B^2 + \dots)] - \frac{\theta_1}{1 - \varphi_1 B} \right]_+ \\ &= \left[ B^{-1} + \frac{\varphi_1}{1 - \varphi_1 B} - \frac{\theta_1}{1 - \varphi_1 B} \right]_+ = \frac{(\varphi_1 - \theta_1)}{1 - \varphi_1 B}, \end{aligned} \tag{1.15}$$

so that

$$\begin{aligned} P_{t-1}X_t &= \frac{(\varphi_1 - \theta_1)}{1 - \varphi_1 B} \left( \frac{1 - \varphi_1 B}{1 - \theta_1 B} \right) X_{t-1} \\ &= \frac{(\varphi_1 - \theta_1)}{1 - \theta_1 B} X_{t-1}; \end{aligned} \tag{1.16}$$

at lag  $\ell$ ,

$$P_{t-\ell}X_t = \left[ \frac{d(B)}{B^\ell} \right]_+ u_{t-\ell} = \left[ \frac{d(B)}{B^\ell} \right]_+ \frac{1}{d(B)} X_{t-\ell} \tag{1.17}$$

where

$$\begin{aligned}
\left[ \frac{d(B)}{B^\ell} \right]_+ &= \left[ B^{-\ell} \frac{(1 - \theta_1 B)}{(1 - \varphi_1 B)} \right]_+ \\
&= \left[ \frac{B^{-\ell}}{1 - \varphi_1 B} - \frac{\theta_1 B^{-(\ell-1)}}{(1 - \varphi_1 B)} \right] \\
&= \left[ \frac{\varphi_1^\ell}{1 - \varphi_1 B} - \frac{\theta_1 \varphi_1^{\ell-1}}{1 - \varphi_1 B} \right] \\
&= \frac{\varphi_1^{\ell-1} (\varphi_1 - \theta_1)}{1 - \varphi_1 B}, \tag{1.18}
\end{aligned}$$

hence

$$\begin{aligned}
P_{t-\ell} X_t &= \frac{(\varphi_1 - \theta_1) \varphi_1^{\ell-1}}{1 - \varphi_1 B} \left( \frac{1 - \varphi_1 B}{1 - \theta_1 B} \right) X_{t-\ell} \\
&= \frac{\varphi_1^{\ell-1} (\varphi_1 - \theta_1)}{1 - \theta_1 B} X_{t-\ell} \tag{1.19}
\end{aligned}$$

or

$$P_t X_{t+\ell} = \frac{\varphi_1^{\ell-1} (\varphi_1 - \theta_1)}{1 - \theta_1 B} X_t.$$

## 2. Chain rule for prediction

Let

$$P_t X_{t+1} = \sum_{j=0}^{\infty} h_j X_{t-j}.$$

Then,

$$\begin{aligned}
P_{t+\ell} X_{t+\ell+1} &= \sum_{j=0}^{\infty} h_j X_{t+\ell-j} \\
&= h_0 X_{t+\ell} + h_1 X_{t+\ell-1} + \cdots + h_\ell X_t + h_{\ell+1} X_{t-1} + \cdots
\end{aligned}$$

$$\begin{aligned}
P_t [P_{t+\ell} X_{t+\ell+1}] &= P_t X_{t+\ell+1} \\
&= h_0 P_t X_{t+\ell} + h_1 P_t X_{t+\ell-1} + \cdots + h_{\ell-1} P_t X_{t+1} \\
&\quad + h_\ell X_t + h_{\ell+1} X_{t-1} + \cdots
\end{aligned}$$



$$\begin{aligned}
&= \sum_{i=0}^{\ell-1} h_i P_t X_{t+\ell-i} + \sum_{i=\ell}^{\infty} h_i X_{t+\ell-i} \\
&= \sum_{i=0}^{\ell-1} h_i P_t X_{t+\ell-i} + \sum_{i=0}^{\infty} h_{i+\ell} X_{t-i} .
\end{aligned}$$

This formula allows one to compute predictions several steps ahead from one-step ahead predictions.

### 3. Properties of prediction errors

Let

$$P_t X_{t+\ell} = P(X_{t+\ell} \mid X_t, X_{t-1}, \dots) .$$

If

$$X_t = \sum_{j=0}^{\infty} d_j u_{t-j}, \quad (3.1)$$

where

$$d_0 = 1, \quad (3.2)$$

$$u_t = X_t - P_{t-1} X_t, \quad (3.3)$$

$$\mathbb{V}(u_t) = \sigma^2, \quad (3.4)$$

then

$$\begin{aligned}
X_{t+\ell} &= \sum_{j=0}^{\infty} d_j u_{t+\ell-j}, \\
P_t X_{t+\ell} &= \sum_{j=\ell}^{\infty} d_j u_{t+\ell-j}.
\end{aligned}$$

$$\begin{aligned}
e_t(\ell) &\equiv X_{t+\ell} - P_t X_{t+\ell} \\
&= \sum_{j=0}^{\ell-1} d_j u_{t+\ell-j} \\
&= u_{t+\ell} + d_1 u_{t+\ell-1} + \dots + d_{\ell-1} u_{t+1}
\end{aligned}$$

follows a  $MA(\ell - 1)$  process, hence

$$\begin{aligned} V[e_t(\ell)] &= \sigma^2 [1 + d_1^2 + \cdots + d_{\ell-1}^2] \\ &= \sigma^2 \sum_{i=0}^{\ell-1} d_i^2, \end{aligned}$$

where  $d_0 = 1$ . Consequently,

$$e_t(1) = u_{t+1}. \quad (3.5)$$

One-step ahead prediction errors associated with optimal are uncorrelated between each others. Further,

$$\begin{aligned} E[e_t(\ell) e_{t+k}(\ell)] &= \sigma^2 \sum_{i=0}^{\ell-k-1} d_i d_{i+k} \\ &= \sigma^2 \sum_{i=k}^{\ell-1} d_i d_{i-k}, \quad 0 \leq k \leq \ell - 1, \end{aligned} \quad (3.6)$$

$$E[e_t(\ell) e_{t+k}(\ell)] = \begin{cases} \sigma^2 \sum_{i=k}^{\ell-1} d_i d_{i-k}, & \text{if } 0 \leq k \leq \ell - 1, \\ 0, & \text{if } k \geq \ell, \end{cases} \quad (3.7)$$

$$\text{Corr}[e_t(\ell), e_{t+k}(\ell)] = \begin{cases} \frac{\sum_{i=k}^{\ell-1} d_i d_{i-k}}{\sum_{j=0}^{\ell-1} d_j^2}, & \text{if } 0 \leq k \leq \ell - 1, \\ 0, & \text{if } k \geq \ell, \end{cases} \quad (3.8)$$

$$\begin{aligned} \text{Cov}[e_t(\ell), e_t(\ell + j)] &= E[e_t(\ell) e_t(\ell + j)] \\ &= \sigma^2 \sum_{i=0}^{\ell-1} d_i d_{i+j} \neq 0, \quad \text{for } j \geq 0. \end{aligned} \quad (3.9)$$

The covariances between prediction errors at different horizons are given by:

$$\begin{aligned} C[e_t(\ell), e_t(\ell + j)] &= E[e_t(\ell) e_t(\ell + j)] \\ &= \sigma^2 \sum_{i=0}^{\ell-1} d_i d_{i+j}. \end{aligned}$$

## 4. Prediction with ARIMA models

Suppose  $X_t$  is an ARIMA process of the form:

$$\varphi_p(B)(1-B)^d X_t = \theta_q(B) u_t + \bar{\mu}, \quad (4.1)$$

$$u_t \sim BB(0, \sigma^2), \quad (4.2)$$

If we denote

$$\varphi(B) = \varphi_p(B)(1-B)^d, \quad (4.3)$$

we can write

$$\varphi(B) X_t = \theta_q(B) u_t + \bar{\mu}, \quad (4.4)$$

or, equivalently,

$$\begin{aligned} & [1 - \varphi_1 B - \dots - \varphi_{p+d} B^{p+d}] X_t \\ & = [1 - \theta_1 B - \dots - \theta_q B^q] u_t + \bar{\mu}, \end{aligned} \quad (4.5)$$

hence the *difference-equation representation* of  $X_t$ :

$$\begin{aligned} X_t & = \varphi_1 X_{t-1} + \dots + \varphi_{p+d} X_{t-p-d} \\ & \quad + u_t - \theta_1 u_{t-1} - \dots - \theta_q u_{t-q} + \bar{\mu} \end{aligned} \quad (4.6)$$

or

$$X_t = \sum_{i=1}^{p+d} \varphi_i X_{t-i} - \sum_{j=1}^q \theta_j u_{t-j} + \bar{\mu} + u_t. \quad (4.7)$$

**4.1 Example** For an ARIMA(1, 1, 0) process, we have:

$$\begin{aligned} (1 - \phi_1 B)(1 - B) X_t & = u_t, \\ [1 - \varphi_1 B + \varphi_2 B^2] X_t & = [1 - (\phi_1 + 1) B + \phi_1 B^2] X_t = u_t, \\ \varphi_1 & = (\phi_1 + 1), \quad \varphi_2 = -\phi_1. \end{aligned}$$

On applying the projection operator  $P_t$  on both sides of the equation (4.7), we obtain:

$$\begin{aligned} P_t X_{t+1} & = \sum_{i=1}^{p+d} \varphi_i P_t X_{t+1-i} - \sum_{j=0}^q \theta_j P_t u_{t+1-j} + \bar{\mu} \\ & = \sum_{i=1}^{p+d} \varphi_i X_{t+1-i} - \sum_{j=1}^q \theta_j u_{t+1-j} + \bar{\mu}, \end{aligned}$$

$$P_t X_{t+2} = \varphi_1 P_t X_{t+1} + \sum_{i=2}^{p+d} \varphi_i X_{t+2-i} - \sum_{j=2}^q \theta_j u_{t+2-j} + \bar{\mu},$$

and, more generally,

$$P_t X_{t+l} = \sum_{i=1}^{\ell-1} \varphi_i P_t X_{t+l-i} + \sum_{i=\ell}^{p+d} \varphi_i X_{t+l-i} - \sum_{j=\ell}^q \theta_j u_{t+l-j} + \bar{\mu}.$$

On noting that

$$P_t u_{t+l} = \begin{cases} 0, & \text{if } l \geq 1, \\ u_{t+l}, & \text{if } l \leq 0, \end{cases}$$

we then see that

$$P_t X_{t+l} = \sum_{i=1}^{p+d} \varphi_i P_t X_{t+l-i} + P_t u_{t+l} - \sum_{j=1}^q \theta_j P_t u_{t+l-j} + \bar{\mu} \quad (4.1)$$

and

$$\begin{aligned} \varphi(B) P_t X_{t+l} &= \theta(B) P_t u_{t+l} + \bar{\mu}, \\ \varphi(B) X_{t+l} &= \theta(B) u_{t+l} + \bar{\mu}, \\ \varphi(B) (X_{t+l} - P_t X_{t+l}) &= \theta(B) (u_{t+l} - P_t u_{t+l}). \end{aligned}$$

Let

$$\begin{aligned} \psi(B) &= 1 + \psi_1 B + \psi_2 B^2 + \dots = \sum_{j=0}^{\infty} \psi_j B^j, \\ \varphi(B) \psi(B) &= \theta(B), \\ e_t(\ell) &= X_{t+l} - P_t X_{t+l}, \quad \ell \geq 1, \\ \varphi(B) e_t(\ell) &= \varphi(B) \psi(B) (u_{t+l} - P_t u_{t+l}). \end{aligned}$$

If we note that

$$\begin{aligned} e_t(\ell) &= 0, \quad \ell \leq 0 \\ u_{t+l} - P_t u_{t+l} &= \begin{cases} 0, & \ell \leq 0 \\ u_{t+l}, & \ell \geq 1 \end{cases} \end{aligned}$$

we can simplify  $\varphi(B)$  on both sides and get

$$\begin{aligned} e_t(\ell) &= \psi(B)(u_{t+\ell} - P_t u_{t+\ell}) \\ &= u_{t+\ell} + \psi_1 u_{t+\ell-1} + \cdots + \psi_{\ell-1} u_{t+1} \\ &= \sum_{i=0}^{\ell-1} \psi_i u_{t+\ell-i}, \end{aligned}$$

where  $\psi_0 = 1$ . It follows that

$$\begin{aligned} \mathbb{E}[e_t(\ell)] &= 0 \\ \mathbb{V}[e_t(\ell)] &= \mathbb{V}(\ell) = \sigma^2 \left[ 1 + \sum_{j=1}^{\ell-1} \psi_j^2 \right]. \end{aligned}$$

One-step ahead prediction errors  $e_t(1)$  are uncorrelated between each other. More generally,

$$\mathbb{E}[e_t(\ell) e_{t-j}(\ell)] = \begin{cases} \sigma^2 \sum_{i=j}^{\ell-1} \psi_i \psi_{i-j}, & \text{if } 0 \leq j \leq \ell - 1, \\ 0, & \text{if } |j| \geq \ell, \end{cases} \quad (4.8)$$

$$\mathbb{E}[e_t(\ell) e_t(\ell + j)] = \left( \sum_{i=0}^{\ell-1} \psi_i \psi_{i+j} \right) \sigma^2.$$

If we assume that  $u_t \stackrel{ind}{\sim} N[0, \sigma_a^2]$

$$e_t(\ell) \sim N \left[ 0, \left\{ 1 + \sum_{j=1}^{\ell-1} \psi_j^2 \right\} \sigma^2 \right],$$

we can compute confidence intervals for predictions:

$$\mathbb{P} [P_t X_{t+\ell} - c_{\alpha/2} \Delta_\ell \leq X_{t+\ell} \leq P_t X_{t+\ell} + c_{\alpha/2} \Delta_\ell] = 1 - \alpha$$

where

$$\Delta_\ell^2 = \left\{ 1 + \sum_{j=1}^{\ell-1} \psi_j^2 \right\} \sigma^2, \quad (4.9)$$

$$\mathbb{P}[N(0, 1) \geq c_\alpha] = \alpha.$$

If we compute predictions at different horizons  $\ell$ , we obtain a *prediction function*:

$$P_t X_{t+\ell} \equiv \hat{X}_t(\ell), \quad \ell = 1, 2, 3, \dots,$$

$$\begin{aligned}
\varphi(B) X_t &= \theta(B) u_t + \bar{\mu}, \\
\varphi(B) \hat{X}_t(\ell) &= \theta(B) P_t u_{t+\ell} + \bar{\mu}, \\
\varphi_p(B) (1-B)^d \hat{X}_t(\ell) &= \theta(B) P_t u_{t+\ell} + \bar{\mu}.
\end{aligned}$$

If  $d = 0$  and  $\ell$  is large, we have:

$$\begin{aligned}
\varphi_p(B) \hat{X}_t(\ell) &\simeq \bar{\mu}, \\
\hat{X}_t(\ell) &\simeq \mu \equiv \frac{\bar{\mu}}{\varphi_p(B)} = \frac{\bar{\mu}}{1 - \varphi_1 - \dots - \varphi_p}.
\end{aligned}$$

If  $d = 1$  and  $\ell$  is large,

$$\begin{aligned}
\varphi_p(B) (1-B) \hat{X}_t(\ell) &\simeq \bar{\mu} \\
(1-B) \hat{X}_t(\ell) &\simeq \mu \equiv \frac{\bar{\mu}}{\varphi_p(B)} \\
\hat{X}_t(\ell) &\simeq \mu_0 + \mu \ell \quad \text{Arithmetic progression.}
\end{aligned}$$

If  $d = 2$  and  $\ell$  is large,

$$\hat{X}_t(\ell) \simeq \mu_0 + \mu_1 \ell + \mu_2 \ell^2.$$

## 5. Bibliographic notes

The reader will find general discussions of prediction based on ARIMA models in Box and Jenkins (1976, Sections 5.1-5.5, 5.7), Brockwell and Davis (1991, Sections 5.1-5.5) and Hamilton (1994, Chap. 4). On prediction for stationary processes and the Wiener-Kolmogorov formula, see also Wiener (1949), Whittle (1983), Whiteman (1983) and Sargent (1987).

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